Change in serum 25-hydroxyvitamin D and parallel change in depressive symptoms in Dutch older adults

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Abstract

Objective: Previous prospective studies on the association between vitamin D status and depression used a single 25-hydroxyvitamin D (25(OH)D) measurement. We investigated the association between change in serum 25(OH)D and parallel change in depressive symptoms over time in Dutch older adults.

Design: A population-based, prospective study in two cohorts of older men and women from the Longitudinal Aging Study Amsterdam.

Methods: Serum 25(OH)D concentrations were determined at two time points: in 1995/1996 and 13 years later in the older cohort (aged 65–88y, n = 173) and in 2002/2003 and 6 years later in the younger cohort (55–65 years, n = 450). At these time points, depressive symptoms were measured with the Center for Epidemiologic Studies Depression scale (CES-D). Associations were tested by multiple linear regression analyses.

Results: During follow-up, serum 25(OH)D concentrations increased in 32.4% of the older cohort and in 69.8% of the younger cohort. In the older cohort, change in 25(OH)D was not associated with change in CES-D score. In the younger cohort, no associations were observed in participants with higher baseline 25(OH)D concentrations (>58.6 nmol/L), but in those with lower baseline 25(OH)D concentrations, an increase in 25(OH)D was associated with a decrease in CES-D score (adjusted B per 10 nmol/L 25(OH)D increase: −0.62 (95% CI: −1.17, −0.07)).

Conclusions: Our study suggests that over 6 years, an increase in serum 25(OH)D is associated with a small decrease in depressive symptoms in young older adults with lower baseline 25(OH)D. Well-designed intervention trials are required to determine causality.

Introduction

Depression is a major cause of disability and common in older adults (1, 2). Clinically relevant depressive symptoms affect 8 to 16% of the community-dwelling older population, and prevalence estimates of major depression in this population range from 1 to 4% (3). Depression has adverse health consequences, such as disability, cardiovascular disease, diabetes and premature death, as well as social and economic consequences (4). However, the aetiology of late-life depression is poorly understood.

Several pathways are involved in the aetiology of late-life depression. Vitamin D, a neurosteroid hormone, may...
play a role in the development of depression, as vitamin D receptors and the activating enzyme 1-α-hydroxylase are present in areas of the human brain that are related to mood (5, 6). In addition, the active metabolite of vitamin D, 1,25-dihydroxyvitamin D, influences the synthesis of monoamine neurotransmitters such as serotonin. This metabolite also affects the regulation of neurotrophic factors and stimulates anti-inflammatory actions in the brain (6, 7, 8). Low serum 25-hydroxyvitamin D (25(OH)D) concentrations are highly prevalent in older adults, as a result of reduced exposure to sunlight, a decreased efficiency of the skin to synthesise vitamin D₃ and lower dietary intake (9).

Previous research has suggested an association between vitamin D status and depression. In many cross-sectional studies, lower 25(OH)D concentrations were associated with a higher risk of depressive symptoms or depression (e.g. (10, 11)), while in others, vitamin D and depression were not associated (12, 13). Prospective studies showed a higher incidence of depressive symptoms or depression over 1–6 years in older persons who had low baseline vitamin D levels (14, 15, 16, 17, 18). However, no prospective associations were found over 4–15 years in older adults (19, 20), older men (21, 22) and adults (23). Similarly, we recently found no cross-sectional and prospective associations of serum 25(OH)D concentrations with depressive symptoms in Dutch older men and women, except for a higher risk of depressive symptoms over 6 years in older women with baseline 25(OH)D concentrations below 75 nmol/L (24). Differences in measures of vitamin D status and depression, study populations, analysis techniques, covariates and latitude may have given rise to these inconsistencies. Furthermore, experimental studies have found inconsistent results as well (e.g. (25, 26)). Meta-analyses of randomised controlled trials (RCTs) did not find support for the efficacy of vitamin D supplementation in depression; however, heterogeneity between the included trials was high (27, 28, 29, 30). A recent review on observational studies and RCTs has indicated that vitamin D deficiency may be a risk factor for late-life depression based on observational data but not on experimental data (31).

Previous research thus shows that the potential causal role of vitamin D status in depression risk is currently not clear. Next to well-designed RCTs, prospective studies are of interest to gain a better insight into this association in the general population. One of the limitations of all previous prospective studies is that only a single measurement of serum 25(OH)D was used (14, 15, 16, 17, 18, 19, 20, 21, 22, 23, 24). Vitamin D levels vary due to changes in sunlight exposure, diet and supplement use (32, 33); therefore, multiple intra-individual measurements are needed for a better understanding of the longitudinal association. To date, no prospective study has addressed the association between changes in serum 25(OH)D concentrations and changes in depressive symptoms over time. Within two cohorts of the Longitudinal Aging Study Amsterdam (LASA), repeated measurements of serum 25(OH)D and depressive symptoms provided the opportunity to prospectively study whether changes in serum 25(OH)D over time are associated with parallel changes in depressive symptoms in Dutch older adults. We hypothesised that an increase in 25(OH)D concentration would lead to a decrease in severity of depressive symptoms.

Subjects and methods

Study population

LASA is an ongoing Dutch cohort study on the determinants, trajectories and consequences of physical, cognitive, emotional and social functioning in relation to ageing. Participants were recruited from three geographic regions in the Netherlands, constituting a nationally representative sample of the Dutch older population. Measurement cycles are carried out every 3 years at the participants’ homes, including a main interview, medical interview and questionnaire. The Medical Ethics Committee of the VU University Medical Center (VUmc) approved the study, and all participants gave informed consent. A detailed description of the sampling and data collection procedures can be found elsewhere (34, 35).

Participants

For the present study, data from two LASA cohorts were used. At LASA’s start in 1992/1993, 3107 subjects aged 55–85 years were enrolled in the first cohort (‘older cohort’). At the second measurement cycle (1995/1996, ‘baseline’ of this cohort), persons aged ≥65 years who participated in the medical interview (n=1509) were asked to donate a blood sample. Serum 25(OH)D concentration was measured in 1320 persons. Serum 25(OH)D was again measured 13 years later (2008/2009, n=390). For this study, persons with 25(OH)D measurements at both cycles were selected (n=175). Two persons were excluded because of missing data on
baseline covariates, resulting in an analytical sample of 173 persons. The recruitment and exclusion of persons from the older cohort are shown in Fig. 1A; attrition was mainly due to mortality (63.0%) (34).

In 2002/2003, a second LASA cohort was recruited consisting of 1002 persons aged 55–65 years (‘younger cohort’). The medical interview was completed by 919 persons. Serum 25(OH)D concentration was measured in 2002/2003 (baseline, \(n=739\)) and 6 years later (2008/2009, \(n=525\)). A total of 458 persons had 25(OH)D measurements at both cycles, of which eight were excluded due to missing data on the Center for Epidemiologic Studies Depression scale (CES-D) at follow-up (\(n=1\)) or on baseline covariates (\(n=7\)). The recruitment and exclusion of persons from the younger cohort are shown in Fig. 1B.

**Depression measurements**

At all measurement cycles, depressive symptoms were assessed by the CES-D, a 20-item self-report scale about depressive symptoms experienced in the previous week (36). The CES-D has been widely used in older population-based samples and is a valid and reliable instrument for Dutch older adults (37). Total scores range from 0 to 60, with a higher score indicating more depressive symptoms. Change in depressive symptoms over 13 or 6 years was calculated by subtracting the CES-D score at baseline from the score at the follow-up cycle.

**25(OH)D measurements**

Morning blood samples were drawn in 1995/1996 (older cohort) and 2002/2003 (younger cohort). Participants...
were allowed to have tea and toast for breakfast, but no dairy products. In 2009 (both cohorts), fasting blood samples were drawn: participants were not allowed to take any food or drink from midnight. All samples were centrifuged and stored at −20°C until 25(OH)D determination in 1997/1998 (1995/1996 samples) and in 2010/2011 (2002/2003 and 2009 samples). A competitive protein-binding assay was used in 1997/1998 (Nichols Diagnostics Capistrano, CA, USA; interassay coefficient of variation (CV): 10%), whereas a radioimmunoassay was used in 2010/2011 (Diasorin, Stillwater, Minnesota, USA; interassay CV: 10%). All measurements were carried out at the Endocrine Laboratory of the VUMc.

Before change in 25(OH)D concentration could be calculated, the 25(OH)D measurements needed to be (1) standardised for assay and (2) corrected for season. First, as part of the 7th Framework Programme project ODIN, all serum 25(OH)D data have been standardised by applying the Vitamin D Standardisation Program protocols (38, 39). Without differences in assay, 25(OH)D change can be properly calculated (30). Second, individual concentrations were deseasonalised before calculating the intra-individual change, as blood samples were drawn throughout the year and 25(OH)D concentrations vary across seasons (highest in summer and autumn, lowest in winter and spring) due to changes in ultraviolet radiation exposure (32, 40). For this, a sinusoidal regression model was fitted:

Original25(OH)Dconcentration = \( \beta_0 + \beta_1 \sin(2\pi T / 365.25) + \beta_2 \cos(2\pi T / 365.25) \)

where original 25(OH)D concentration denotes the standardised 25(OH)D concentration, \( T \) denotes the day of the year of blood draw and \( \beta_i \) (\( i = 0, 1, 2 \)) are estimated regression coefficients. To obtain individual deseasonalised 25(OH)D concentrations, the residuals (difference between the predicted value and the original value) from this model were added to the annual mean of the model (the intercept: \( \beta_0 \)) (41, 42). In this way, the seasonal variation of 25(OH)D concentration was adjusted for. Deseasonalising was performed separately for each cohort and each measurement cycle: for 1995/1996 and 2009 for the older cohort and for 2002/2003 and 2009 for the younger cohort. Subsequently, intra-individual change in 25(OH)D concentration was calculated by subtracting the standardised, deseasonalised 25(OH)D concentration at baseline from the standardised, deseasonalised 25(OH)D concentration at follow-up (13 or 6 years later). Unless otherwise specified, standardised, deseasonalised 25(OH)D data were used and reported in this study.

Other variables

Potential confounders were examined at baseline and included sex, age, education level, BMI, smoking status, alcohol use, physical activity and number of chronic diseases. Data on sex and age were derived from municipal registries. Participants were asked about their highest completed level of education, and this was converted into education years. BMI was calculated as measured body weight (in kg) divided by squared measured body height (in m). Self-reported smoking status was categorised into never, former and current smoking. Categories of self-reported alcohol use were no, light, moderate and (very) excessive (43). Physical activity during the past 2 weeks (walking outdoors, cycling, sports and heavy household activities, in min/day) was assessed using the validated LASA Physical Activity Questionnaire (44). Self-reported number of chronic diseases included seven major somatic diseases: asthma/chronic obstructive pulmonary disease, cardiac disease, peripheral arterial disease, diabetes mellitus, cerebrovascular accident/stroke, osteoarthritis/rheumatoid arthritis and cancer. Use of antidepressants (yes/no, only used in sensitivity analyses) was retrieved during the medical interview using ATC codes of the medications that were used at that time by the participant.

Statistical analyses

As age range and follow-up time between the two cycles differed between the two cohorts, all analyses were conducted separately for each cohort. Absolute change in 25(OH)D concentrations was calculated over 13 years (older cohort) and over 6 years (younger cohort). Absolute change in CES-D score was also calculated over these time periods and checked for normality. Paired-samples \( t \)-tests were performed to assess whether changes over time in 25(OH)D and CES-D were statistically significant. Serum 25(OH)D change was used as continuous determinant, expressed per 10nmol/L. To compare participants with relatively large 25(OH)D decreases with those with large increases, it was also divided in tertiles with tertile 3 (increase in 25(OH)D concentration) serving as reference category. For the continuous baseline characteristics, linear regression analyses were used to examine linear trends across the tertiles of 25(OH)D change. For the categorical variables, chi-square (linear-by-linear association) tests were performed to test for trend.
The associations between change in 25(OH)D and change in depressive symptoms over 13 or 6 years were analysed by multivariable linear regression analysis. Based on previous studies on 25(OH)D and depression (11, 14, 15, 24), sex was examined as a potential effect modifier. Another potential effect modifier was baseline serum 25(OH)D concentration, as this may influence the change in serum 25(OH)D (45). Effect modification was tested by including sex or baseline serum 25(OH)D (dichotomous, median-split) and the interaction term of the potential effect modifier with 25(OH)D change in the unadjusted regression models. A P value <0.10 of the interaction term(s) was considered statistically significant. Three models were made: model 1 was adjusted for baseline serum 25(OH)D (to interpret the regression coefficient for subjects with the same 25(OH)D concentration at baseline); model 2 was additionally adjusted for age, sex and education level; and model 3 was additionally adjusted for the lifestyle factors BMI, physical activity, alcohol use and smoking. To test linear trends across the tertiles, the categorical tertile variable was used as a continuous variable to retrieve a P value for trend.

Three sensitivity analyses were conducted. First, all analyses were additionally adjusted for number of chronic diseases in a separate model, as this covariate might confound and/or mediate the association causing potential overadjustment. Second, in order to check whether changes in lifestyle factors affect the association, model 3 was additionally adjusted for the changes in BMI, physical activity (both absolute continuous change), alcohol use and smoking (both categorical: stable, increase, decrease). Third, all analyses were repeated after exclusion of participants who used antidepressants at baseline or at one or more of the measurement cycles during the follow-up period.

All analyses were performed using SPSS version 23 (SPSS Inc.). A double-sided P value of <0.05 was considered statistically significant.

**Results**

**Baseline characteristics**

Participants from the older cohort (n=173) had a median age of 68.9 years (interquartile range (IQR): 67.0–72.3) and the mean baseline deseasonalised serum 25(OH)D was 58.2 nmol/L (standard deviation (s.d.): 16.1) (Table 1). The median age of participants from the younger cohort (n=450) was 59.8 years (IQR: 57.0–62.2) and their mean baseline deseasonalised serum 25(OH)D was 59.5 nmol/L. (s.d.: 17.9). When comparing the 25(OH)D change tertiles from the older cohort, participants from tertile 1 tended to be more educated and tended to drink more alcohol compared to the other tertiles. In the younger cohort, participants from tertile 1 were more often women and tended to have more chronic diseases.

Comparison of included (n=1336, Fig. 1A) participants from the older cohort showed that excluded participants had lower 25(OH)D concentrations (P<0.001), more depressive symptoms (P<0.001), were older (P<0.001), less physically active (P<0.001), more often smokers (P=0.002), drank less alcohol (P<0.001) and had more chronic diseases (P<0.001). In the younger cohort, this comparison showed that excluded participants (n=469, Fig. 1B) had more depressive symptoms (P=0.044), were more often smokers (P=0.004) and had more chronic diseases (P=0.019) compared to included participants (n=450).

**Absolute change in serum 25(OH)D concentrations and depressive symptoms**

Figure 2 shows the absolute, intra-individual change in serum 25(OH)D over 13 years for the older cohort and over 6 years for the younger cohort. The difference in distribution between the two cohorts becomes clear from the cutoff values of the tertiles. In the older cohort, 117 participants (67.6%) experienced a decrease in 25(OH)D concentration, whereas this was 172 (38.2%) in the younger cohort. Mean 25(OH)D of participants from the older cohort decreased by 6.9 nmol/L (s.d.: 20.9, P<0.001). The mean CES-D score increased during the same time period by 2.0 points (s.d.: 6.5, P<0.001), suggesting increasing depressive symptoms over time. In the younger cohort, mean 25(OH)D increased by 5.5 nmol/L (s.d.: 17.8, P<0.001), and the mean CES-D score decreased by 0.7 points (s.d.: 6.5) over 6 years (P<0.027). Change in depressive symptoms was normally distributed in both cohorts.

**Association between change in serum 25(OH)D and change in depressive symptoms**

In the older cohort, change in serum 25(OH)D (as continuous determinant) was not associated with change in CES-D score after adjustment for baseline 25(OH)D concentration. Similarly, tertiles of 25(OH)D change did not show an association, and the P for trend was not statistically significant (0.246). No interactions with sex and baseline 25(OH)D were observed.
Table 1  Baseline characteristics of two cohorts of older adults and according to tertiles of change in serum 25(OH)D. Values are displayed as mean±s.d. or median (interquartile range), and as %.

<table>
<thead>
<tr>
<th></th>
<th>Older cohort (65–88 years in 1995/1996)</th>
<th>All</th>
<th>Tertile 1&lt;sup&gt;b&lt;/sup&gt;</th>
<th>Tertile 2&lt;sup&gt;b&lt;/sup&gt;</th>
<th>Tertile 3&lt;sup&gt;b&lt;/sup&gt;</th>
<th>P for trend</th>
<th>Younger cohort (55–65 years in 2002/2003)</th>
<th>All</th>
<th>Tertile 1&lt;sup&gt;c&lt;/sup&gt;</th>
<th>Tertile 2&lt;sup&gt;c&lt;/sup&gt;</th>
<th>Tertile 3&lt;sup&gt;c&lt;/sup&gt;</th>
<th>P for trend</th>
</tr>
</thead>
<tbody>
<tr>
<td>N</td>
<td>173</td>
<td></td>
<td>58</td>
<td>58</td>
<td>57</td>
<td>0.333</td>
<td>450</td>
<td></td>
<td>150</td>
<td>150</td>
<td>150</td>
<td>0.983</td>
</tr>
<tr>
<td>Age (years)</td>
<td>68.9 (67.0–72.3)</td>
<td></td>
<td>69.5 (67.0–74.2)</td>
<td>69.0 (67.1–72.2)</td>
<td>68.6 (66.9–71.6)</td>
<td>0.333</td>
<td>59.8 (57.0–62.2)</td>
<td></td>
<td>59.2 (57.1–62.2)</td>
<td>60.2 (56.7–62.3)</td>
<td>59.7 (57.0–62.2)</td>
<td>0.983</td>
</tr>
<tr>
<td>Sex (% women)</td>
<td>54.9</td>
<td></td>
<td>48.3</td>
<td>56.9</td>
<td>59.6</td>
<td>0.221</td>
<td>53.1</td>
<td></td>
<td>10.0 (9.0–11.0)</td>
<td>10.0 (9.0–11.0)</td>
<td>10.0 (9.0–12.8)</td>
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<tr>
<td>Education (years)</td>
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<td>9.0 (6.0–10.3)</td>
<td>9.0 (6.0–11.0)</td>
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<td>10.0 (9.0–11.0)</td>
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<td>10.0 (9.0–11.0)</td>
<td>10.0 (9.0–11.0)</td>
<td>10.0 (9.0–12.8)</td>
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</tr>
<tr>
<td>BMI (kg/m&lt;sup&gt;2&lt;/sup&gt;)&lt;sup&gt;d&lt;/sup&gt;</td>
<td>26.3 (24.2–29.0)</td>
<td></td>
<td>25.0 (23.7–29.7)</td>
<td>26.8 (24.8–28.4)</td>
<td>26.7 (23.9–29.5)</td>
<td>0.599</td>
<td>26.8 (24.4–29.5)</td>
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<td>27.3 (24.1–30.4)</td>
<td>26.6 (24.1–28.9)</td>
<td>26.9 (24.6–29.6)</td>
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<tr>
<td>Physical activity (min/day)&lt;sup&gt;d&lt;/sup&gt;</td>
<td>67.2 (33.2–107.4)</td>
<td></td>
<td>76.1 (32.7–117.1)</td>
<td>51.6 (28.6–92.9)</td>
<td>76.4 (43.4–115.8)</td>
<td>0.623</td>
<td>66.7 (38.5–115.7)</td>
<td></td>
<td>63.8 (37.7–108.0)</td>
<td>72.1 (39.4–112.5)</td>
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<td>Alcohol use (%)</td>
<td></td>
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<td>6.9</td>
<td>17.2</td>
<td>14.0</td>
<td>0.057</td>
<td>6.7</td>
<td></td>
<td>10.0</td>
<td>5.3</td>
<td>4.7</td>
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<td>58.6</td>
<td>53.4</td>
<td>46.7</td>
<td>0.057</td>
<td>48.7</td>
<td></td>
<td>45.3</td>
<td>50.0</td>
<td>50.7</td>
<td>0.690</td>
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<tr>
<td>Moderate</td>
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<td>24.1</td>
<td>20.7</td>
<td>14.0</td>
<td>0.057</td>
<td>35.6</td>
<td></td>
<td>35.3</td>
<td>36.0</td>
<td>35.3</td>
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</tr>
<tr>
<td>(Very) excessive</td>
<td>8.1</td>
<td></td>
<td>10.3</td>
<td>8.6</td>
<td>5.3</td>
<td>0.057</td>
<td>9.1</td>
<td></td>
<td>9.3</td>
<td>8.7</td>
<td>9.3</td>
<td>0.690</td>
</tr>
<tr>
<td>Smoking (%)</td>
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<td></td>
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<td>0.807</td>
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<tr>
<td>Never</td>
<td>36.4</td>
<td></td>
<td>36.2</td>
<td>32.8</td>
<td>40.4</td>
<td>0.057</td>
<td>24.9</td>
<td></td>
<td>26.0</td>
<td>24.7</td>
<td>24.7</td>
<td>0.871</td>
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<td>Former</td>
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<td>55.2</td>
<td>56.9</td>
<td>49.1</td>
<td>0.057</td>
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<td>54.7</td>
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<td>0.871</td>
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<tr>
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<td></td>
<td>8.6</td>
<td>10.3</td>
<td>10.5</td>
<td>0.057</td>
<td>25.1</td>
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<td>27.3</td>
<td>20.7</td>
<td>27.3</td>
<td>0.871</td>
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<td>Number of chronic diseases</td>
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<td></td>
<td>1.0 (0.0–1.0)</td>
<td>1.0 (0.0–2.0)</td>
<td>1.0 (0.0–1.0)</td>
<td>0.301</td>
<td>1.0 (0.0–1.0)</td>
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<td>0.0 (0.0–1.0)</td>
<td>1.0 (0.0–1.0)</td>
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<tr>
<td>Serum 25(OH)D (nmol/L)</td>
<td>58.2 ± 16.4</td>
<td></td>
<td>67.3 ± 15.0</td>
<td>54.7 ± 14.6</td>
<td>52.4 ± 15.9</td>
<td>&lt;0.001</td>
<td>56.1 ± 18.2</td>
<td></td>
<td>63.6 ± 19.4</td>
<td>53.2 ± 16.8</td>
<td>51.4 ± 15.8</td>
<td>&lt;0.001</td>
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<td>Season of blood collection (% winter)</td>
<td>50.9</td>
<td></td>
<td>50.0</td>
<td>46.6</td>
<td>56.1</td>
<td>0.514</td>
<td>71.6</td>
<td></td>
<td>64.7</td>
<td>71.3</td>
<td>78.7</td>
<td>0.007</td>
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<tr>
<td>Deseasonalised serum 25(OH)D (nmol/L)</td>
<td>58.2 ± 16.1</td>
<td></td>
<td>67.3 ± 15.0</td>
<td>54.4 ± 13.6</td>
<td>52.8 ± 15.8</td>
<td>&lt;0.001</td>
<td>59.5 ± 17.9</td>
<td></td>
<td>66.4 ± 19.3</td>
<td>57.1 ± 16.6</td>
<td>54.9 ± 15.6</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Depressive symptoms (CES-D score)&lt;sup&gt;d&lt;/sup&gt;</td>
<td>4.0 (1.0–9.0)</td>
<td></td>
<td>3.0 (1.8–8.0)</td>
<td>5.0 (1.8–10.3)</td>
<td>4.0 (1.0–10.5)</td>
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<td>5.0 (2.0–10.0)</td>
<td>5.0 (2.0–11.0)</td>
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<td>Use of antidepressants (%)</td>
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<td>3.4</td>
<td>3.4</td>
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<td>0.221</td>
<td>3.8</td>
<td></td>
<td>3.3</td>
<td>4.7</td>
<td>3.3</td>
<td>1.000</td>
</tr>
</tbody>
</table>

<sup>a</sup>The follow-up time was 13 years for the older cohort and 6 years for the younger cohort. <sup>b</sup>Tertiles older cohort: T1 −78.6 to −16.4, T2 −16.3 to −0.7, T3 −0.6 to 66.2 nmol/L. <sup>c</sup>Tertiles younger cohort: T1 −54.0 to −2.7, T2 −2.7 to 11.8, T3 11.9 to 81.3 nmol/L. <sup>d</sup>Trends across the 25(OH)D change tertiles were tested with linear regression analyses after natural logtransformation of these variables. 25(OH)D, 25-hydroxyvitamin D; BMI, body mass index; CES-D, Center for Epidemiologic Studies Depression Scale; T, tertile.
In this study, we observed no associations between change in serum 25(OH)D concentration and parallel change in depressive symptoms in a cohort of older adults (65–88 years at baseline, follow-up 13 years). Neither did we find associations in a younger cohort (55–65 years at baseline, follow-up 6 years) with a baseline 25(OH)D concentration above 59 nmol/L. However, in participants from the younger cohort with a lower baseline 25(OH)D, we observed that an increase in 25(OH)D concentration was associated with a decrease in severity of depressive symptoms.

The observed associations between change in 25(OH)D concentration and change in depressive symptoms differed between the two cohorts. In the younger cohort, a 10 nmol/L increase in serum 25(OH)D was associated with a 0.62 points lower CES-D score, but only if baseline 25(OH)D concentrations were relatively low. The null findings in the younger cohort’s group with relatively
Table 2  Associations between change in deseasonalised serum 25(OH)D and change in CES-D score.

<table>
<thead>
<tr>
<th></th>
<th>Older cohort, 13 years follow-up (n = 173)</th>
<th>Younger cohort, 6 years follow-up: Baseline deseasonalised 25(OH)D &lt;58.6 nmol/L (n = 225)</th>
<th>Younger cohort, 6 years follow-up: Baseline deseasonalised 25(OH)D &gt;58.6 nmol/L (n = 225)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Model 1</td>
<td>Model 2</td>
<td>Model 3</td>
</tr>
<tr>
<td></td>
<td>n</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Change in 25(OH)D</td>
<td>-0.27</td>
<td>-0.22</td>
<td>-0.16</td>
</tr>
<tr>
<td>Change in 25(OH)D in tertiles</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Tertile 1</td>
<td>58</td>
<td>0.25</td>
<td>0.63</td>
</tr>
<tr>
<td>Tertile 2</td>
<td>58</td>
<td>0.05</td>
<td>0.64</td>
</tr>
<tr>
<td>Tertile 3</td>
<td>57</td>
<td>Reference</td>
<td>Reference</td>
</tr>
<tr>
<td>P for trend</td>
<td>0.103</td>
<td>0.162</td>
<td>0.246</td>
</tr>
<tr>
<td>Change in 25(OH)D</td>
<td>-0.63</td>
<td>-0.64</td>
<td>-0.62</td>
</tr>
<tr>
<td>Change in 25(OH)D in tertiles</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Tertile 1</td>
<td>51</td>
<td>3.30</td>
<td>3.07</td>
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<td>Tertile 2</td>
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<td>0.005</td>
<td>0.008</td>
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<tr>
<td>Change in 25(OH)D</td>
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<td>0.23</td>
<td>0.27</td>
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<td>Change in 25(OH)D in tertiles</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Tertile 1</td>
<td>99</td>
<td>-0.80</td>
<td>-0.85</td>
</tr>
<tr>
<td>Tertile 2</td>
<td>68</td>
<td>-1.12</td>
<td>-0.92</td>
</tr>
<tr>
<td>Tertile 3</td>
<td>58</td>
<td>Reference</td>
<td>Reference</td>
</tr>
<tr>
<td>P for trend</td>
<td>0.501</td>
<td>0.451</td>
<td>0.470</td>
</tr>
</tbody>
</table>

Model 1: adjusted for baseline standardised and deseasonalised 25(OH)D concentration. Model 2: additionally adjusted for age, sex and education level. Model 3: additionally adjusted for BMI, physical activity, alcohol use and smoking.

*P-value <0.05, **P-value <0.01; †Change in 25(OH)D (continuous, per 10 nmol/l). P for trend; †, unstandardised regression coefficient.

High baseline 25(OH)D concentrations suggest that a change in 25(OH)D concentration above a certain level – we used the median of 59 nmol/L – has less influence on depressive symptoms. However, this was not confirmed in the older cohort. An explanation for the non-significant association in the older cohort may be the small sample size (n = 173), which limited the statistical power. The betas of the older cohort were in the same direction and of comparable size as those in the younger cohort’s group with relatively low baseline 25(OH)D concentrations.

To the best of our knowledge, this is the first observational study that investigated intra-individual change in 25(OH)D in relation to intra-individual change in depressive symptoms. Our results suggest that supplementation of vitamin D in those with relatively low baseline levels might positively influence depressive symptoms. Indeed, a 12-month vitamin D supplementation trial in overweight adults showed that 25(OH)D concentrations increased from 55 to 112 nmol/L in the group receiving 40 000 IU vitamin D₃ per week and from 52 to 88 nmol/L in the group receiving 20 000 IU per week, compared to no change (from 52 to 50 nmol/L) in the placebo group; the treatment groups had a significant greater reduction in depressive symptoms than the placebo group (25). However, other RCTs, which reported pre- and post-intervention 25(OH)D concentrations, showed a larger 25(OH)D increase in the treatment group but no effect on depression (26, 46) or mental well-being (47, 48). Thus, more research is needed to further elucidate the potential impact of 25(OH)D change on depressive symptoms.

In our previous prospective study, an inverse association between baseline 25(OH)D and depressive symptoms over 6 years was found only in women from the older LASA cohort (24). In the present study, with change in 25(OH)D as determinant instead of baseline 25(OH)D, sex was not an effect modifier in the older cohort. Explanations for this might be differences in the study sample and the limited power (n = 173 vs 1282). In addition, De Koning, et al. (24) found no association in the younger cohort, whereas we did for those with low baseline 25(OH)D. This discrepancy might be explained by general health status: being healthy might be important to withstand negative effects on mood of a low 25(OH)D, but not negative effects of a decrease in 25(OH)D.
In the present study, we also observed a difference between the two cohorts with respect to the distributions of 25(OH)D changes: 62% of the younger cohort had an increase in serum 25(OH)D concentration over 6 years, compared to 32% of the older cohort over 13 years. This difference might be explained by age differences (55–65 years in the younger cohort vs 65–88 years in the older cohort). Older persons may have reduced their outdoor activities with ageing due to for example functional limitations, leading to less exposure to sunlight (9, 49). Our results are similar to those of the Tromsø Study, in which subjects older than 65 years at baseline had a decrease in serum 25(OH)D of 0.3 nmol/L over 14 years, while those younger than 65 years had an increase of 2.0 nmol/L (P<0.05) (50). Furthermore, our mean 13-year change in 25(OH)D in the older cohort (−6.9 nmol/L) and 6-year change in the younger cohort (+5.5 nmol/L) were comparable but slightly different to previous findings within LASA that were based on different statistical models and on measurements that were not yet standardised. Van Schoor et al. (40) showed that 25(OH)D concentrations decreased with 6.5 nmol/L in the older cohort and increased with 2.4 nmol/L in the younger cohort (see their Table 3: age per year (longitudinal)).

Our study has several strengths, including its longitudinal design, a nationally representative sample of older adults, and the use of deseasonalised serum 25(OH)D values at two time points. Further, the measurements of 25(OH)D were standardised for assay. Limitations of this study are the relatively healthy study population – as most participants needed to visit the hospital for blood sampling – and the exclusion of participants with only one serum 25(OH)D measurement that resulted in smaller and less representative samples. Moreover, our study was limited by the use of self-reported depressive symptoms during the past week instead of a clinical assessment; however, the CES-D scale has been shown to have good psychometric properties in samples of older adults (37). Reverse causality is a limitation as well, since we do not know when the changes occurred during the 13 or 6 years; an increase in depressive symptoms might have occurred before the decrease in 25(OH)D. However, our previous paper showed that baseline 25(OH)D was associated with the onset of depressive symptoms over 6 years in older women without depressive symptoms at baseline (24). This indicated an association in the direction from 25(OH)D to depression, which was the assumption for the present study. Lastly, no data on supplement use or nutritional intake of vitamin D were available. However, measured 25(OH)D concentrations are a more valid estimate of vitamin D status than intake from foods and supplements (33).

In conclusion, our study suggests that a decrease in serum 25(OH)D concentration over 6 years is associated with a parallel increase in depressive symptoms in young older adults with a low baseline 25(OH)D. We did not observe an association in young older adults with a higher baseline 25(OH)D over 6 years, or in older adults over a period of 13 years. Whether specific groups may emotionally benefit from an increase in vitamin D concentration (e.g. persons with relative low 25(OH)D concentrations) needs further investigation; well-designed RCTs are needed to determine the potential causal association between vitamin D and depression.

Declaration of interest
The authors declare that there is no conflict of interest that could be perceived as prejudicing the impartiality of this study.

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