Thyroid antibodies and risk of preterm delivery: a meta-analysis of prospective cohort studies

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Abstract

Background: Observational studies suggest possible associations between thyroid antibodies and risk of preterm delivery. However, whether thyroid antibodies are risk factors of preterm labor remains controversial. Our goal was to evaluate the associations between thyroid antibodies and risk of preterm delivery by conducting a meta-analysis of prospective cohort studies.

Methods: PubMed, Embase, and Wangfang databases were searched through January 2012 to identify studies that met pre-stated inclusion criteria. Data were extracted using standardized forms. Either a fixed- or a random-effects model was used to calculate the overall combined relative ratio (RR) with its corresponding 95% confidence interval (95% CI) to evaluate the relationship between thyroid antibodies and preterm delivery risk. Subgroup analyses were mainly performed by type of thyroid antibodies including thyroid peroxidase antibody (TPO-Ab) and thyroglobulin antibody (TG-Ab).

Results: Eleven prospective cohort studies involving 35 467 participants were included. The combined RR of preterm delivery for pregnant women with thyroid antibodies compared with the reference group was 1.41 (95% CI 1.08–1.84, P=0.011). Subgroup analysis yielded the combined RR of preterm delivery for pregnant women with TPO-Ab compared with the reference group was 1.69 (95% CI 1.19–2.41, P=0.003), whereas pregnant women with positive TG-Ab had no obvious risk of preterm delivery compared with the reference group (RR=0.88, 95% CI 0.60–1.29, P=0.513). Sensitivity analysis restricted to studies excluding women with thyroid dysfunction yielded similar results. Meta-regression analysis suggested that the status of exclusion or inclusion of women with thyroid dysfunction was the major source of heterogeneity in this meta-analysis. No evidence of publication bias was observed.

Conclusions: Current evidence suggests that the presence of TPO-Ab in pregnant women significantly increases the risk of preterm delivery.

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Introduction

Preterm delivery is defined as birth occurring at or before 37 weeks of gestation and it remains one of the most intractable problems that contributes to perinatal morbidity and mortality in obstetric practice (1, 2). Despite a concerted effort to decrease the incidence of preterm delivery, there has been a relentless increased incidence over the last two decades (2). Preterm delivery results from a series of disorders, implicating maternal and fetal disease, some of which are explained and interrelated and others of which are of unknown cause (1, 2). Research efforts to address this problem have risen substantially over the past 10 years but have not resulted in improvements in prediction and prevention of preterm delivery (2). The incidence of preterm delivery is associated with various epidemiological and clinical risk factors including a previous preterm birth, periodontal disease, low maternal BMI, smoking, history of induced abortion, anemia, assisted reproduction, multifetal gestation, and other established risk factors, and a good understanding of risk factors may lead to the development of new therapeutic strategies (2, 3, 4, 5, 6). Though there have been great advances in our understanding of the molecular and cellular pathways operative in reproductive tissues in the maintenance of uterine quiescence during pregnancy, and in initiating term and preterm labor, our knowledge of the pathophysiology of preterm delivery is limited (7). Thus, more risk factors of preterm delivery should be identified by epidemiological studies to provide good evidence for the prevention and treatment of preterm delivery (2, 3).
Many epidemiological studies have investigated the links between thyroid antibodies and risk of preterm delivery (8, 9, 10, 11, 12, 13, 14). However, the magnitudes of the association varied between studies and whether thyroid antibodies were risk factors of preterm delivery remains controversial (8, 9, 10, 11, 12, 13, 14). Although a meta-analysis by Thangaratinam et al. (15) combined several cohort studies and reported a statistically significant relation of thyroid antibodies to preterm delivery risk, the evidence was limited because only five cohort studies were available at that time. Besides, Thangaratinam et al. (15) neither performed meta-analyses by the type of thyroid antibodies including thyroid peroxidase antibody (TPO-Ab) and thyroglobulin antibody (TG-Ab) nor explored the possible sources of obvious heterogeneity. Furthermore, whether thyroid antibodies are risk factors or merely silent markers of preterm delivery remains unclear. An improved understanding of this issue may have important public health and clinical implications given the possibility that prevention and treatment of preterm delivery might be well performed and the incidence of preterm delivery can be effectively reduced in pregnant women with thyroid antibodies. Therefore, with recently accumulating evidence, our goal was to evaluate the associations between thyroid antibodies (both TPO-Ab and TG-Ab) and risk of preterm delivery by conducting a meta-analysis of prospective cohort studies.

**Materials and methods**

**Search strategy and study selection**

We attempted to follow the proposed Meta-Analysis of Observational Studies in Epidemiology (MOOSE) guidelines to report the present meta-analysis (16). We conducted a literature search in PubMed, Embase, and Wangfang databases through January 2012 for relevant studies that tested the associations between thyroid antibodies (TPO-Ab or TG-Ab) and risk of preterm delivery. The following search terms were used: i) thyroid autoimmune antibody, thyroid auto-immune antibodies, thyroid antibody, thyroid antibodies, thyroid autoantibody, thyroid autoantibodies, thyroid globulin antibody, thyroid globulin antibodies, TPO-Ab, or thyroid peroxidase antibodies; ii) preterm delivery, preterm labor, preterm birth, premature delivery, premature labor, or premature birth; and iii) cohort study, cohort studies, prospective study, prospective studies, follow-up study, or follow-up studies. No restrictions were imposed. In addition, we reviewed the reference lists of retrieved papers and recent reviews. The reference lists of included studies were also hand searched. When necessary, additional information was sought from the authors. We first performed an initial screening of titles or abstracts. A second screening was based on full-text review. Studies were considered eligible if they met the following criteria: i) the study design was a prospective cohort study; ii) the exposure of interest was thyroid antibodies (TPO-Ab or TG-Ab); iii) the outcome of interest was preterm delivery; and iv) relative risk (RR) and the corresponding 95% confidence interval (95% CI; or data to calculate them) were reported. Data from abstracts, review articles, editorials, case reports, and letters were not included. Studies performed in assisted reproductive technology (ART) pregnancies were also excluded owing to the high risk of preterm delivery and the sample size from those studies. Overlapping study or studies containing overlapping participants were excluded.

**Data extraction**

The key exposure variable was the presence or absence of thyroid antibodies (TPO-Ab or TG-Ab) at baseline. In all studies, women without thyroid antibodies (TPO-Ab or TG-Ab) served as the reference group. Outcome of interest in this study was preterm delivery. Preterm delivery was defined as birth occurring at or before 37 weeks of gestation (1). Data extraction was then performed using a standardized data collection form. We extracted any reported RR or incidence density ratios of outcomes for patients with thyroid antibodies (TPO-Ab or TG-Ab) compared with the reference group. We also extracted study characteristics for each trial and contacted primary authors to get additional information on the studies if needed. Data were recorded as follows: first author’s name, year of publication, country of origin, study period, characteristics of study population and age at baseline, number of preterm delivery events and total participants, potential confounders, and ascertainment of thyroid antibodies. Two authors independently conducted the studies’ selection and data extraction. Any disagreements were resolved by discussion. Information was extracted from each selected article on study characteristics, quality, and test results.

**Assessment of quality of methods**

We used the Newcastle–Ottawa scale to assess methodological quality of the selected studies, with the components of study design that were related to internal validity (17, 18). Information on adequacy of definition of cohorts, representativeness of the sample, selection and evaluation of controls, comparability, ascertainment of exposure, and outcome were evaluated for cohort studies (17, 18). The study was considered to have low risk of bias if it scored a maximum of four for selection, two for comparability, and three for assessment of outcome or ascertainment of exposure (17, 18). Any study that scored 1 or zero for selection or zero for comparability or for outcome assessment was categorized as having a high risk of bias. Studies that scored in between were rated as having a medium risk of bias.
Statistical analyses

RR with its corresponding 95% CI was used as a common measure of the association between thyroid antibodies (TPO-Ab or TG-Ab) and risk of preterm delivery. We calculated the overall combined RR with its 95% CI to assess the strength of the relationships between thyroid antibodies and preterm delivery risk. The significance of the combined RR was determined by the Z test and a P value of <0.05 was considered significant. In our study, two models of meta-analysis for dichotomous outcomes were conducted: the random-effects model and the fixed-effects model (19, 20). The random-effects model was conducted using the DerSimonian and Laird’s method, which assumed that studies were taken from populations with varying effect sizes and calculated the study weights both from in-study and between-study variances (20). The fixed-effects model was conducted using the Mantel–Haenszel’s method, which assumed that studies were sampled from populations with the same effect size and made an adjustment to the study weights according to the in-study variance (19). To assess the between-study heterogeneity more precisely, both the χ²-based Q statistic test (Cochran’s Q statistic) to test for heterogeneity and the I² statistic to quantify the proportion of the total variation due to heterogeneity were calculated (21, 22). The I² index expressing the percentage of the total variation across studies due to heterogeneity was calculated to assess the between-study heterogeneity. I² values of 25, 50, and 75% were used as evidence of low, moderate, and high heterogeneity respectively (21). If moderate or high heterogeneity existed, the random-effects model was used to pool the results; otherwise, the fixed-effects model was used to pool the results when I² value was < 50%. Because characteristics of participants and ascertainment of thyroid antibodies were not consistent between studies, we subsequently conducted meta-regression analysis to explore possible explanations for heterogeneity, and a P value > 0.1 was selected to indicate absence of significant heterogeneity of the estimates (23). We conducted a meta-regression analysis to explore predefined sources of heterogeneity of our primary predictor of interest. The following prespecified variables were analyzed: proportion of thyroid antibodies, multivariate adjustment (yes/no), study quality, publication year, sample size, and the status of exclusion or inclusion of women with thyroid dysfunction. Multivariate meta-regression analyses were also undertaken. Subgroup meta-analyses were performed by the type of thyroid antibodies (TPO-Ab or TG-Ab) and the status of exclusion or inclusion of women with thyroid dysfunction. We also investigated the influence of a single study on the overall risk estimate by omitting one study in each turn to validate the credibility of outcomes in this meta-analysis (24). Potential publication bias was assessed by visual inspection of the Begg’s funnel plots, in which the standard error of logor of each study was plotted against its logor and an asymmetric plot suggested possible publication bias (25). In addition, we also performed the Egger linear regression test at the P<0.10 level of significance to assess the funnel-plot’s asymmetry (26).

All analyses were performed using STATA version 12.0 (StataCorp LP, College Station, TX, USA). A P value <0.05 was considered statistically significant, except where otherwise specified.

Results

Study selection and study characteristics

We initially retrieved 33 unique citations from the PubMed, Embase, and Wangfang databases. Of these, 16 citations were excluded after the first screening based on the abstracts or titles, mainly because they were overlapping records, not cohort studies, or obviously irrelevant studies. After full-text review of the remaining 17 articles (8, 9, 10, 11, 12, 13, 14, 27, 28, 29, 30, 31, 32, 33, 34, 35, 36), three articles were excluded because the exposure was not relevant (32, 33, 36), two articles were excluded because they were not on preterm delivery (34, 35), and one article was excluded because of the nested case–control design (31). Finally, 11 articles involving a total of 35,467 participants were included in our meta-analysis (8, 9, 10, 11, 12, 13, 14, 27, 28, 29, 30). A flow chart showing the study selection was presented in Fig. 1.

The characteristics of the 11 prospective cohort studies were presented in Table 1. These cohort studies were published between 1990 and 2011. Four studies
Table 1 Characteristics of the 11 prospective cohort studies.

<table>
<thead>
<tr>
<th>References</th>
<th>Location</th>
<th>Period</th>
<th>Population (events of preterm delivery)</th>
<th>Definition of thyroid dysfunction</th>
<th>Confounders</th>
<th>Exclusion criteria</th>
<th>Thyroid antibody (positive percent, %)</th>
</tr>
</thead>
<tbody>
<tr>
<td>(28) USA</td>
<td>NG</td>
<td></td>
<td>552 pregnant women recruited before 13th week of pregnancy from seven obstetric practice groups (124)</td>
<td>NG</td>
<td>None</td>
<td>None</td>
<td>TG-Ab or TPO-Ab positive (14.9%)</td>
</tr>
<tr>
<td>(14) Belgium</td>
<td>June 1990 to December 1992</td>
<td></td>
<td>1660 pregnant women recruited from the prenatal clinic (62)</td>
<td>Total T&lt;sub&gt;4&lt;/sub&gt; and T&lt;sub&gt;3&lt;/sub&gt; values out of the reference ranges (adjusted for gestational time), abnormal free T&lt;sub&gt;4&lt;/sub&gt; index, and serum TSH out of the reference ranges 0.20–4 mIU/l</td>
<td>None</td>
<td>Thyroid dysfunction</td>
<td>TPO-Ab positive (12.8%)</td>
</tr>
<tr>
<td>(13) Japan</td>
<td>NG</td>
<td></td>
<td>1179 healthy pregnant women who attended maternity clinic (43)</td>
<td>NG</td>
<td>None</td>
<td>Multi-fetal gestations</td>
<td>TPO-Ab positive (11.2%); TG-Ab positive (3.3%)</td>
</tr>
<tr>
<td>(11) Italy</td>
<td>November 2002 to October 2004</td>
<td></td>
<td>984 pregnant women who attended the department of obstetrics (84)</td>
<td>Serum TSH out of the reference range 0.27–4.2 mIU/l; FT&lt;sub&gt;4&lt;/sub&gt; out of the reference range 9.3–18.0 ng/l (12–33.5 pmol/l); FT&lt;sub&gt;3&lt;/sub&gt; out of the reference range 0.27–4.2 mIU/l</td>
<td>None</td>
<td>Thyroid dysfunction</td>
<td>TPO-Ab positive (6.3%)</td>
</tr>
<tr>
<td>(12) Pakistan</td>
<td>NG</td>
<td></td>
<td>1500 pregnant women registered for antenatal care at obstetrics department (152)</td>
<td>Serum TSH out of the reference range 0.27–4.2 mIU/l</td>
<td>None</td>
<td>Thyroid dysfunction</td>
<td>TPO-Ab positive (11.2%)</td>
</tr>
<tr>
<td>(30) China</td>
<td>June 2006 to February 2008</td>
<td></td>
<td>864 pregnant women who attended the department of obstetrics (45)</td>
<td>Serum TSH out of the reference range 0.25–4.0 mIU/l; or FT&lt;sub&gt;3&lt;/sub&gt; &lt; 2.5 pmol/l; or FT&lt;sub&gt;4&lt;/sub&gt; &lt; 11.5 pmol/l</td>
<td>None</td>
<td>Thyroid dysfunction</td>
<td>TPO-Ab positive (13.0%)</td>
</tr>
<tr>
<td>(10) Finland</td>
<td>July 1985 to June 1986</td>
<td></td>
<td>9247 singleton pregnancies (312)</td>
<td>Mothers with serum concentrations of both TSH and FT&lt;sub&gt;4&lt;/sub&gt; between the 5th and 95th percentiles were considered to have normal thyroid function</td>
<td>None</td>
<td>Multiple pregnancy</td>
<td>TPO-Ab positive (5.0%); TG-Ab positive (5.0%)</td>
</tr>
<tr>
<td>(9) USA</td>
<td>October 1999 to December 2002</td>
<td></td>
<td>10 062 singleton pregnancies (663)</td>
<td>NG</td>
<td>Age, BMI, gravidity, race, smoking, gestational diabetes, educational level, and recruitment site</td>
<td>Multiple pregnancy</td>
<td>TG-Ab or TPO-Ab positive (14.6%)</td>
</tr>
<tr>
<td>(8) USA</td>
<td>2005</td>
<td></td>
<td>4516 euthyroid women in the first trimester (217)</td>
<td>NG</td>
<td>None</td>
<td>Thyroid dysfunction or miscarriages</td>
<td>TPO-Ab positive (6.8%)</td>
</tr>
<tr>
<td>(29) India</td>
<td>January 2007 to April 2007</td>
<td></td>
<td>483 consecutive pregnant women in the first trimester (22)</td>
<td>TSH &gt; 4 mIU/l</td>
<td>None</td>
<td>Thyroid dysfunction</td>
<td>TPO-Ab positive (9.3%)</td>
</tr>
<tr>
<td>(27) Europe</td>
<td>March 2006 to December 2006</td>
<td></td>
<td>4420 singleton pregnancies with no history of thyroid disease (102)</td>
<td>NG</td>
<td>None</td>
<td>Preeclampsia</td>
<td>TPO-Ab positive (10.2%); TG-Ab positive (13.6%)</td>
</tr>
</tbody>
</table>

NG, data were not given in the original study; TPO-Ab, thyroid peroxidase antibody; TG-Ab, thyroglobulin antibody.
were conducted in Europe (10, 11, 14, 27), three in the United States (8, 9, 28), and four in Asia (12, 13, 29, 30). The sizes of the cohorts ranged from 438 to 10 062 (total 35 467). The ascertainment of thyroid antibodies varied across studies, with most based on TPO-Ab positive. Among the 11 articles included here, nine evaluated TPO-Ab (8, 10, 11, 12, 13, 14, 27, 28, 29, 30), three evaluated TG-Ab (10, 13, 27), and two mixed TG-Ab and TPO-Ab (9, 28). Three articles reported data on both TPO-Ab and TG-Ab, thus were extracted as six individual studies (10, 13, 27). The exclusion criteria also varied across studies, with six studies excluding women with thyroid dysfunction (8, 11, 12, 14, 29, 30) and five studies including women with thyroid dysfunction (9, 10, 13, 27, 28). The TPO-Ab positive rate ranged from 5.3 to 14.9%, while the TG-Ab positive rate ranged from 3.4 to 15.8%. Six studies reported information on the definition of thyroid dysfunction (10, 11, 12, 14, 29, 30) and only one study reported information regarding potential confounders (age, BMI, gravidity, race, smoking, gestational diabetes, educational level, and recruitment site) (9). In addition, there was an obvious difference in the definition of thyroid dysfunction in those studies included in this meta-analysis (Table 1). All 11 cohort studies had low risk of bias for selection and outcome assessment on the Newcastle–Ottawa scale (14).

**Synthesis analysis**

Summary of the pooled outcomes in this meta-analysis was shown in Table 2.

**Thyroid antibodies and preterm delivery** There were 14 individual data from 11 articles including nine on TPO-Ab, three on TG-Ab, and two on TG-Ab and TPO-Ab (8, 9, 10, 11, 12, 13, 14, 27, 28, 29, 30). Meta-analysis showed that the overall combined RR of preterm delivery risk for pregnant women with thyroid antibodies compared with the reference group was 1.41 (95% CI 1.08–1.84, \( P = 0.011 \); Fig. 2). Sensitivity analyses by sequential omission of individual studies did not materially alter the overall combined RR, suggesting that the combined RR was valid and credible.

**TPO-Ab and preterm delivery** Nine prospective cohort studies reported relevant data on the association between TPO-Ab and preterm delivery risk (8, 10, 11, 12, 13, 14, 27, 29, 30). Meta-analysis showed that the combined RR of preterm delivery risk for pregnant women with positive TPO-Ab compared with the reference group was 1.69 (95% CI 1.19–2.41, \( P = 0.003 \); Fig. 3). Sensitivity analyses by sequential omission of individual studies did not materially alter the overall combined RR, suggesting that the combined RR was valid and credible.

**TG-Ab and preterm delivery** Three prospective cohort studies reported relevant data on the association between TG-Ab and preterm delivery risk (10, 13, 27). Meta-analysis showed that pregnant women with positive TG-Ab had no obvious risk of preterm delivery compared with the reference group (RR = 0.88, 95% CI 0.60–1.29, \( P = 0.513 \); Fig. 4). Sensitivity analyses by sequential omission of individual studies did not materially alter the overall combined RR, suggesting that the combined RR was valid and credible.

**Heterogeneity analysis**

There was obvious heterogeneity for the meta-analyses of both all 11 studies and the subgroup analysis of TPO-Ab (\( I^2 = 75.0 \) and 71.6% respectively; Table 2). Meta-regression analyses suggested that the status of exclusion or inclusion of women with thyroid dysfunction was the major source of heterogeneity in this meta-analysis (for total 11 studies, \( P = 0.001 \); for subgroup analysis of TPO-Ab, \( P = 0.029 \); however, other sources were not found by meta-regression analyses. Multivariate analyses of meta-regression were also performed but did not find further definite source of heterogeneity. For the meta-analysis of all 11 studies, the status of thyroid function could explain 63.57% of between-study variance (\( P = 0.001 \); for subgroup analysis of TPO-Ab, \( P = 0.029 \) ) with residual heterogeneity (residual \( I^2 = 58.94\% \) ).

As the status of exclusion or inclusion of women with thyroid dysfunction was the major source of heterogeneity, sensitivity analyses were further performed to assess the influence of the status of exclusion or inclusion of women with thyroid dysfunction on the associations between thyroid antibodies and preterm delivery risk. In studies excluding women with thyroid dysfunction, the overall combined RR of preterm delivery risk for pregnant women with thyroid antibodies or with TPO-Ab positive status was significant (for thyroid antibodies, \( RR = 1.98, 95\% CI 1.29–3.04, P = 0.002 \); for TPO-Ab, \( RR = 1.98, 95\% CI 1.29–3.04, P = 0.002 \); Table 2). In studies including women with thyroid dysfunction, none of the overall combined RRs for pregnant women with thyroid antibodies, TPO-Ab positive status, and TG-Ab positive status were statistically significant (for thyroid antibodies, \( RR = 1.12, 95\% CI 0.97–1.29, P = 0.118 \); for TPO-Ab, \( RR = 1.22, 95\% CI 0.87–1.71, P = 0.243 \); for TG-Ab, \( RR = 0.88, 95\% CI 0.60–1.29, P = 0.513 \); Table 2). Subgroup analysis of women with TG-Ab positive status was not performed owing to the lack of relevant studies.

**Publication bias**

Visual inspection of the Begg’s funnel plot did not identify substantial asymmetry (Supplementary Figure 1, see section on supplementary data given at www.eje-online.org
the end of this article). The Egger linear regression test also showed no evidence of publication bias among studies of thyroid antibodies and risk of preterm delivery (Egger test, $P = 0.990$). Thus, the publication bias was not evident in this meta-analysis.

Discussion

Many epidemiological studies have examined the associations between thyroid antibodies and the risk of preterm delivery, but the magnitudes of the associations vary between studies and whether thyroid antibodies are risk factors of preterm delivery remains controversial (8, 9, 10, 11, 12, 13, 14). Although a meta-analysis combined several cohort studies and reported a statistically significant relation of thyroid antibodies to preterm delivery risk, evidence was limited because only five cohort studies were available at that time and there was unexplained heterogeneity in that meta-analysis (15). Several new studies have emerged since that meta-analysis was completed (8, 27, 29, 30), and there was a need for a new meta-analysis to comprehensively assess the associations between thyroid antibodies (TPO-Ab or TG-Ab) and risk of preterm delivery and to find some possible explanations for the heterogeneity in the previous meta-analysis (15). Therefore, we performed this updated meta-analysis by including 11 prospective cohort studies involving 35,467 participants (8, 9, 10, 11, 12, 13, 14, 27, 28, 29, 30). The combined RR of preterm delivery for pregnant women with thyroid antibodies compared with the reference group was 1.41 (95% CI 1.08–1.84, $P = 0.011$). Subgroup analysis showed that the combined RR of preterm delivery for pregnant women with positive TPO-Ab compared with the reference group was 1.69 (95% CI 1.19–2.41, $P = 0.003$), whereas pregnant women with positive TG-Ab had no obvious risk of preterm delivery compared with the reference group (RR = 0.88, 95% CI 0.60–1.29, $P = 0.513$). There was obvious heterogeneity for the meta-analysis of both all 11 studies and the subgroup analysis of TPO-Ab ($I^2 = 75.0$ and 71.6% respectively; Table 2). Meta-regression analyses suggested that the status of exclusion or inclusion of women with thyroid dysfunction was the major source of heterogeneity in this meta-analysis (for all 11 studies, $P = 0.001$; for subgroup analysis of TPO-Ab, $P = 0.029$). Besides, sensitivity analyses restricted to studies excluding women with thyroid dysfunction yielded similar results. Thus, current evidence suggests that pregnant women with positive TPO-Ab have higher risk of preterm delivery.

Heterogeneity is a potential problem when interpreting the results of all meta-analyses, and finding the

<table>
<thead>
<tr>
<th>Analysis items</th>
<th>Studies (participants)</th>
<th>RR (95% CI)</th>
<th>$P_{RR}$</th>
<th>Model</th>
<th>$I^2$ (%)</th>
<th>Absolute effect</th>
<th>Egger test</th>
</tr>
</thead>
<tbody>
<tr>
<td>Thyroid antibodies</td>
<td>14 (40,609)</td>
<td>1.41 (1.08–1.84)</td>
<td>0.011</td>
<td>Random</td>
<td>75.0</td>
<td>22 more per 1000 (from 4 more to 45 more)</td>
<td>0.990</td>
</tr>
<tr>
<td>Studies with euthyroid women</td>
<td>6 (7,932)</td>
<td>1.98 (1.29–3.04)</td>
<td>0.002</td>
<td>Random</td>
<td>71.9</td>
<td>65 more per 1000 (from 19 more to 136 more)</td>
<td>0.209</td>
</tr>
<tr>
<td>Studies including thyroid dysfunction</td>
<td>8 (32,677)</td>
<td>1.12 (0.97–1.29)</td>
<td>0.118</td>
<td>Fixed</td>
<td>0.0</td>
<td>6 more per 1000 (from 2 fewer to 15 more)</td>
<td>0.441</td>
</tr>
<tr>
<td>TPO-Ab</td>
<td>9 (19,191)</td>
<td>1.69 (1.19–2.41)</td>
<td>0.003</td>
<td>Random</td>
<td>71.6</td>
<td>22 more per 1000 (from 4 more to 45 more)</td>
<td>0.134</td>
</tr>
<tr>
<td>Studies with euthyroid women</td>
<td>6 (7,932)</td>
<td>1.98 (1.29–3.04)</td>
<td>0.002</td>
<td>Random</td>
<td>71.9</td>
<td>65 more per 1000 (from 19 more to 136 more)</td>
<td>0.209</td>
</tr>
<tr>
<td>Studies including thyroid dysfunction</td>
<td>3 (11,259)</td>
<td>1.22 (0.87–1.71)</td>
<td>0.243</td>
<td>Fixed</td>
<td>0.0</td>
<td>6 more per 1000 (from 2 fewer to 15 more)</td>
<td>0.734</td>
</tr>
<tr>
<td>TG-Ab</td>
<td>3 (11,108)</td>
<td>0.88 (0.60–1.29)</td>
<td>0.513</td>
<td>Fixed</td>
<td>0.0</td>
<td>5 fewer per 1000 (from 16 fewer to 12 more)</td>
<td>0.916</td>
</tr>
<tr>
<td>Studies including thyroid dysfunction</td>
<td>3 (11,108)</td>
<td>0.88 (0.60–1.29)</td>
<td>0.513</td>
<td>Fixed</td>
<td>0.0</td>
<td>5 fewer per 1000 (from 16 fewer to 12 more)</td>
<td>0.916</td>
</tr>
</tbody>
</table>

TPO-Ab, thyroid peroxidase antibody; TG-Ab, thyroglobulin antibody; RR, risk ratio; 95% CI, 95% confidence interval; random, random-effects model; fixed, fixed-effects model.
The presence of TPO-Ab can reflect a dysregulated activity of the immune system at the fetal–maternal interface and might be a marker of immune dysfunction (43). Thus, plausible biological explanations exist to support the association between TPO-Ab and preterm delivery risk.

A major strength of our study is that all the included original studies use a prospective cohort design, which eliminates the possibility of reverse causation and minimizes selection bias. Moreover, the association of TPO-Ab with risk of preterm delivery persists and remains statistically significant in sensitivity analyses restricted to studies excluding women with thyroid dysfunction. In addition, with the accumulating evidence, enlarged sample size, and elimination of bias caused by thyroid function, we have enhanced statistical power to provide more precise and reliable risk estimates. The prevalence of TPO-Ab ranged from 5.26 to 14.89%, and the overall combined RR of preterm delivery risk for euthyroid pregnant women with TPO-Ab compared with the reference group was 1.98 (95% CI 1.29–3.04, P = 0.002). This improved understanding of the association between TPO-Ab and preterm delivery risk may have important public health and clinical implications given the possibility that prevention of preterm delivery in pregnant women with TPO-Ab might be effectively achieved (1). Negro et al. reported that euthyroid pregnant women who are positive for TPOAb have an increased risk of miscarriage and premature deliveries, and substitutive treatment with l-T4 is able to lower the chance of miscarriage and premature delivery in women with TPOAb positive status (11). Thus, for women with TPOAb positive status, the chance of premature delivery could be lowered by preventive therapy. Thus, TPO-Ab could be a risk factor of preterm delivery and TPO-Ab testing may be a good marker for the prevention of preterm delivery (2, 3).

Several limitations should be taken into account in the interpretation of results from this meta-analysis. First, there was obvious difference in the test assay platforms for TPO-Ab and thresholds for test positivity used between studies, which could result in the

![Figure 2](image-url)
heterogeneity in this meta-analysis. Besides, definition of thyroid dysfunction in pregnancy has evolved over the years and there was obvious difference in the definition of thyroid dysfunction in those studies included in the present meta-analysis. The differences in both the definition of thyroid dysfunction in pregnancy and the threshold for test positivity of TPO-Ab or TG-Ab may cause the heterogeneity and should be taken into account in the interpretation of results. Further studies may identify the more reasonable assay platform or threshold for test positivity to better predict the risk of preterm delivery. Secondly, our meta-analysis used unadjusted RRs from individual cohorts owing to the lack of adjusted RRs. It is no doubt RRs adjusted by potential confounders (smoking, age, history of induced abortion, history of preterm delivery, multifetal gestation, and other established risk factors) are suitable to get a more precise estimate. Further studies can investigate the effects of thyroid antibodies on preterm delivery risk by adjusting the potential confounders. Besides, the differences in the potential confounders prove to be helpful in the heterogeneity exploration, but owing to the limited reported information regarding potential confounders in the included studies, we were unable to identify whether those differences were a source of heterogeneity in this meta-analysis. Thirdly, there was a lack of data on the association between TG-Ab and preterm delivery risk in euthyroid women. Though subgroup meta-analysis of three studies suggested that TG-Ab positive pregnant women had no obvious risk of preterm delivery compared with the reference group (RR = 0.88, 95% CI 0.60–1.29, \( P=0.513 \)), those three did not exclude women with thyroid dysfunction and the accurate estimate might be biased by thyroid dysfunction (10, 13, 27). Besides, this estimate might be an over-interpretation of findings, given that the analysis is based on the results of the quantitative synthesis of only three studies and probably underpowered. Thus, further studies are needed to assess the possible association between TG-Ab and preterm delivery risk in euthyroid women.

Fourthly, there is a broad range of ‘TSH normality’ and obvious difference in the TSH levels in women with TPO-Ab and women without TPO-Ab even in studies excluding women with thyroid dysfunction. Among those six studies excluding women with thyroid dysfunction (8, 11, 12, 14, 29, 30), only Negro et al. compared the TSH levels in women with TPO-Ab with those without TPO-Ab and reported higher TSH levels in women with positive TPO-Ab (1.7 ± 0.5 mIU/l) compared with women without TPO-Ab (1.1 ± 0.4 mIU/l, \( P<0.05 \)) (11), and the difference of TSH levels might bias the real estimate. However, none of those studies reported RRs adjusted by TSH levels, and further studies are needed to assess RRs adjusted by TSH levels, which may further identify whether TPO-Ab is a risk factor independent of conventional risk factors of preterm delivery. Finally, the associations between other types of antibodies and preterm delivery risk are interesting and are worth investigating, such as antinuclear antibody, anti-DNA antibody, and antimitochondrial antibody (13). However, there is little literature assessing the association between other types of antibodies and preterm delivery risk, and we are unable to perform a meta-analysis owing to the limited studies. More studies are needed to assess the possible associations between other types of antibodies and preterm delivery risk in the future. Besides, there is no literature published to assess the association between TPO-Ab and preterm delivery risk in ART pregnancies up to now. Further studies are needed to assess the possible association in ART pregnancies.

This meta-analysis of prospective cohort studies suggests that TPO-Ab in pregnant women significantly increases the risk of preterm delivery. Further studies can investigate the effects of other thyroid antibodies on preterm delivery risk by adjusting the potential confounders.

Supplementary data

This is linked to the online version of the paper at http://dx.doi.org/10.1530/EJE-12-0379.

Declaration of interest

The authors declare that there is no conflict of interest that could be perceived as prejudicing the impartiality of the research reported.

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